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Assessing Host-Virus Codivergence for Close Relatives of Merkel Cell Polyomavirus Infecting African Great Apes

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ABSTRACT

It has long been hypothesized that polyomaviruses (PyV; family *Polyomaviridae*) codiverged with their animal hosts. In contrast, recent analyses suggested that codivergence may only marginally influence the evolution of PyV. We reassess this question by focusing on a single lineage of PyV infecting hominine hosts, the Merkel cell polyomavirus (MCPyV) lineage. By characterizing the genetic diversity of these viruses in seven African great ape taxa, we show that they exhibit very strong host specificity. Reconciliation analyses identify more codivergence than noncodivergence events. In addition, we find that a number of host and PyV divergence events are synchronous. Collectively, our results support codivergence as the dominant process at play during the evolution of the MCPyV lineage. More generally, our results add to the growing body of evidence suggesting an ancient and stable association of PyV and their animal hosts.

IMPORTANCE

The processes involved in viral evolution and the interaction of viruses with their hosts are of great scientific interest and public health relevance. It has long been thought that the genetic diversity of double-stranded DNA viruses was generated over long periods of time, similar to typical host evolutionary timescales. This was also hypothesized for polyomaviruses (family *Polyomaviridae*), a group comprising several human pathogens, but this remains a point of controversy. Here, we investigate this question by focusing on a single lineage of polyomaviruses that infect both humans and their closest relatives, the African great apes. We show that these viruses exhibit considerable host specificity and that their evolution largely mirrors that of their hosts, suggesting that codivergence with their hosts played a major role in their diversification. Our results provide statistical evidence in favor of an association of polyomaviruses and their hosts over millions of years.

Viral diversification is notably shaped by processes that promote host specificity, for example, antagonistic coevolution (1), and opportunities to colonize new hosts, i.e., cross-species transmission events. Depending on their balance, host-virus codivergence patterns may arise and persist over the long term. Long-term codivergence may have played an important role in the diversification of some double-stranded DNA (dsDNA) viruses, e.g., herpesviruses and papillomaviruses (2–5).

Polyomaviruses (PyV; family *Polyomaviridae*) are small nonenveloped viruses with a circular double-stranded DNA genome (ca. 5 kb in length) (6). They infect a broad range of animals, including arthropods and vertebrates (fish, birds, and mammals), and comprise at least 13 distinct viruses infecting humans (7, 8). In humans, infections occur in childhood, persist lifelong, and are usually asymptomatic (9). At least five PyV have been associated with disease in immunosuppressed individuals (10–12). Routes of transmission are poorly characterized but may involve respiratory droplets and/or environmental contamination.

Putative codivergence events of hosts and their PyV have repeatedly been invoked in the literature to explain the structure of PyV diversity. Reconciliation analyses performed on the family scale sometimes supported a significant contribution of codivergence events (8, 13), but other studies have failed to detect any global codivergence signal (14, 15). Similarly, authors focusing on more recent evolutionary events defended opposing views as to the potential codivergence of humans and JC polyomaviruses (JCV) (16, 17–19). An alternative scenario combining ancient noncodivergence events and subsequent lineage-specific codivergence with their hosts, as proposed for papillomaviruses (3), still remains to be tested. The disparate sampling of the PyV animal hosts as well as the lack of resolution of many internal branches of this viral family tree severely compromises the power to detect such patterns from currently available data.

To overcome these limitations, we designed a formal test to assess the influence of codivergence on the evolution of PyV and characterized the genetic diversity of a single lineage of PyV that infects a set of recently diverged host species with a well-resolved phylogeny. Specifically, we focused on viruses infecting African great apes (here, referred to simply as great apes) belonging to the lineage comprising the Merkel cell polyomavirus (MCPyV), an oncogenic human virus (*Human polyomavirus* 5, genus *Alphapolyomavirus*) (10, 20, 21, 22).

MATERIALS AND METHODS

Samples. We collected a total of 386 fecal samples in the wild from seven great ape taxa (Table 1). Great ape samples were collected opportunistically or from habituated animals and preserved in RNAlater (Qiagen, Hilden, Germany), in liquid nitrogen, or by drying over silica. We also collected 197 fecal samples from two human populations in Côte d'Ivoire and the Democratic Republic of the Congo (Table 1). Human samples were preserved in liquid nitrogen. For animal samples, authorization was obtained from responsible local authorities. For human samples, institutional authorization was received along with the written consent of all participants in the study.

Molecular biology. DNA extraction was performed using a Roboklon stool kit (Roboklon, Berlin, Germany), according to manufacturer's instructions.

To identify Merkel cell polyomavirus (MCPyV)-related sequences in DNA extracts, a nested PCR assay was set up that made use of generic, degenerate primers targeting a ca. 700-bp VP1 fragment (Table 2, PCR1). These primers were designed on the basis of published MCPyV sequences and those of MCPyV-related PyV of nonhuman primates (NHP). First-round PCR mixes were set up so as to reduce the risk of carryover contamination with PCR products. The mixtures contained 0.2 μ M each primer, 200 μ M deoxynucleoside triphosphate (dNTP) mix (with dUTP replacing dTTP), 0.3 U of AmpErase uracil *N*-glycosylase (UNG; Invitrogen, Carlsbad, CA, USA), 4 mM MgCl₂, 1× PCR buffer, and 1.25 U of Platinum Taq polymerase (Invitrogen). Second-round PCR mixes were prepared in the same way but did not include UNG. Cycling conditions were as follows: 7 min at 45°C (UNG activity) and 7 min at 95°C, followed

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by 47 cycles (first round) or 45 cycles (second round) of 30 s at 95°C, 30 s at 57°C (first round) or 58°C (second round), and 2 min at 72°C, with a final 10 min at 72°C.

Twenty-two positive samples were then selected based on the results of preliminary phylogenetic analyses to attempt additional nested long-distance (LD) amplification of partial genomes (approximately 2.5 kb) with generic, degenerate primers (Table 2, PCR2) using a TaKaRa-Ex kit (TaKaRa Bio, Inc., Otsu, Japan) according to the manufacturer's instructions. Nondegenerate primers (sequences available from the authors upon request) were used for amplification of the remaining part (approximately 2.8 kb) of the genome with LD nested PCR. LD PCR cycling conditions followed those reported in Leendertz et al. (21).

A total of 174 human DNA extracts were also screened using a seminested PCR system targeting a ca. 200-bp VP1 fragment (Table 2, PCR3). This system was designed to be specific to members of lineage 1 (see below) and was validated on a selection of great ape DNA extracts of known status before being employed on human DNA extracts (data not shown). PCR mix preparation and cycling conditions followed those mentioned above.

Short PCR products were purified using ExoSAP-IT (Affymetrix, Santa Clara, CA, USA) whereas LD PCR products were purified using a column-based PCR purification kit (Qiagen, Venlo, Netherlands). All purified products were sequenced with a BigDye Terminator cycle sequencing kit on a 377 DNA automated sequencer (Applied Biosystems, Warrington, United Kingdom).

Overlapping partial sequences were used to reconstruct circular genomes using Geneious, version 7.1.4 (Biomatters, Ltd., Auckland, New Zealand) (23). Genomes were subsequently annotated with Geneious.

Phylogenetic analyses. Partial VP1 and complete genome data sets were assembled that comprised sequences generated in this study and a selection of (partial VP1) or all (complete genome) MCPyV sequences as well as any publicly available great ape MCPyV-related sequences. Both data sets were reduced to unique sequences and aligned using MUSCLE, as implemented in SeaView, version 4 (24). Conserved nucleotide blocks were selected from the alignments using Gblocks (in SeaView) (25) and used for recombination analyses using RDP4, version 4.46 (26). Two final alignments were generated: one with 74 sequences and 838 positions and another with 16 sequences and 5,150 positions. Further analyses were performed only on the partial VP1 alignment as this comprised more genetic diversity.

The best model of nucleotide substitution (general time-reversible matrix with rate variation across sites [GTR+G₄]) was selected with jModelTest, version 2.1.4 (27), using the Bayesian information criterion. Maximum likelihood (ML) analyses were performed using PhyML, version 3 (28), as implemented on the PhyML Web server (29). The best-fit root of the ML tree was identified using TempEst, version 1.5 (30; http: //tree.bio.ed.ac.uk/software/tempest/). Bayesian Markov chain Monte Carlo (BMCMC) analyses were performed in BEAST, version 1.8.2, under a log-normal relaxed clock (uncorrelated) and three different models of diversification: a pure coalescent model assuming a constant population size, a multispecies coalescent model using the 14-species scheme suggested by species delineation analyses (see below), and a birth-death speciation model (31, 32). Convergence of BMCMC runs (at least two runs per model) and appropriate sampling of the posterior were checked with Tracer, version 1.6 (http://tree.bio.ed.ac.uk/software/tracer/). Branch robustness was assessed through nonparametric bootstrapping (250 pseudoreplicates for ML analyses) or posterior probabilities (BMCMC analyses).

Host specificity analyses. Host specificity was assessed by running Bayesian tip-association significance testing (BaTS) on all posterior samples of trees (PST) generated by BMCMC analyses (33). BaTS allows for tests of the correlation of trait states with ancestry while accounting for phylogenetic uncertainty suggested by the PST. It compares observations to a null distribution generated under the assumption that trait values are not influenced by ancestry. Host species/subspecies was defined as the

Species and subspecies	Country	Site	No. of samples	No. of positives	Proportion (% [95% CI]) ^a	within host subspecies (%) ^a	available sequence (% [GenBank accession no. and host subspecies]) ^b
Gorilla gorilla gorilla	Cameroon	Belgique	20	, 1 ,			
		Campo Ma'an National Park	18	0			
		Mambele	19	0			
	Central African Republic	Dzanga-Sangha Special Reserve	23	0			
	Gabon	Loango National Park	25	1	1.9(0.3-7.4)	98.5	99 (HQ385752, Gorilla gorilla gorilla)
Gorilla gorilla diehli	Cameroon	Greater Takamanda-Mone Landscape	22	0	0(0-18.5)	NA	NA
Gorilla beringei beringei	Rwanda	Volcanoes National Park	51	0			
	Uganda	Bwindi Impenetrable National Park	30	-	1.2(0-7.6)	NA	98 (HQ385752, Gorilla gorilla gorilla)
Gorilla beringei graueri	Democratic Republic of	Kahuzi-Biega National Park	34	7	20.6 (9.3-38.4)	74.7	99 (HQ385752, Gorilla gorilla gorilla)
	the Congo						83 (HQ385747, Pan troglodytes verus)
Pan troglodytes troglodytes	Cameroon	Belgique	J	1			
		Cameroun	10	1			
		Campo Ma'an National Park	1	0			
		Mambele	9	1			
	Gabon	Loango National Park	27	3	11.5 (4.8–24.1)	77	95 (HQ385748, Pan troglodytes verus) 94 (HQ385747, Pan troglodytes verus)
Pan troglodytes schweinfurthii	Uganda	Budongo Central Forest Reserve	33	9			
		Kibale Forest National Park	33	11	30.3 (20-43)	76.5	95 (HQ385747, Pan
							troglodytes verus) 94 (HQ385748, Pan
Pan paniscus	Democratic Republic of	Salonga National Park	26	14	53.8 (33.7-72.9)	77.4	troglodytes verus) 91 (HQ385751, Pan
	the Congo						troglodytes verus) 91 (HQ385746, Pan
Homo sapiens	Côte d'Ivoire	Taï National Park	96	16			troglodytes verus)
	Democratic Republic of the Congo	Salonga National Park	101	14	15.2 (10.7–21.2)	66	100 (JF812999, Homo sapiens)
^{<i>a</i>} At the species/subspecies level. NA, not assessed; CI, confidence interval. ^{<i>b</i>} PyV sequences from Western chimpanzees (<i>Pan troglodytes verus</i>) were already available from a previous study (14).	A, not assessed; CI, confidence in	terval.					

TABLE 1 Samples and screening results

TABLE 2 Primers used in this study

PCR	Primer name	Primer sequence (5'–3')	Annealing temp (°C)	Fragment size (kb)
PCR1	PCR1.1-f	TGTGCTCCTAAGCCBGGATG	57	
	PCR1.1-r	ACTACTGGGTATGGRTTYTTMACC		
	PCR1.2-f	CTGAATCCAAGRATGGGAGT	58	0.7
	PCR1.2-r	CATGAAANGCCATTTTNCCACT		
PCR2	PCR2.1-f	CTGAAGYCTGGGACGMTGAG	57	
	PCR2.1-r	GCAAACATRTGRTAATTGACTCCC		
	PCR2.2-f	TCAGACWCCSAGTCCAGAGG	58	2.5
	PCR2.2-r	GCAAATCYARRGGYTCTCCTC		
PCR3	PCR3.1-f	TGATATGCAGCCMAATMWWCARC	58	
	PCR3.1-r	AAACATGTGATAATTGACTCCCTC		
	PCR3.1-f	TGATATGCAGCCMAATMWWCARC	58	0.2
	PCR3.2-r	AATTGACTCCCTCAATAGGAATG		

trait of interest. Its association with ancestry was assessed at the host subspecies level (8 states) and species level (5 states) independently by running separate BaTS analyses during which 500 null replicates per tree were generated. Global as well as state-specific statistics of association were computed (global, association index [AI] and Fitch parsimony score [PS]; state-specific, maximum exclusive single-state clade size [MC]).

To investigate the association of host and PyV diversification processes, we performed PyV species delineation analyses with the R package SPLITS (34), using the maximum clade credibility tree derived from BMCMC analyses performed under the (coalescent) constant population size model. SPLITS implements general mixed Yule coalescent (GMYC) models (34, 35) which are optimized and compared to the null hypothesis that the tree was generated by pure coalescent processes, i.e., that it reflects diversity within a single species. When the GMYC model outperforms the null model, the parts of the tree most likely to have been generated by between-species and within-species processes can be identified, thereby delineating species (according to the phylogenetic species concept).

Codivergence analyses. The degree of topological congruence and the number of events necessary to explain (reconcile) incongruences were assessed using Jane, version 4(36). Jane implements a genetic algorithm to quickly identify the most parsimonious scenarios of coevolution involving several types of events (codivergence, duplication, duplication with host switch, loss, and failure to diverge). As input, it requires host and parasite phylogenies and the respective tip mapping as well as an event cost matrix. A simplified version of the PyV phylogeny was used as input, whereby single-host clades were collapsed. Three sets of costs were tested: (i) set 1 with the parameter codivergence set at 0, duplication at 1 (under the assumption that duplication incurs costs related to within-host speciation, e.g., maintaining of distinct lineages in the face of within-host competition or tropism change within the same host), duplication with host switch at 1 (host switch incurs costs), loss at 1 (prevalence was always high), and failure to diverge at 1 (given their respective evolutionary timescales, viruses are unlikely to fail to diverge when their hosts do so); (ii) set 2, with the same parameters as set 1 but with loss set at 0 (prevalence may have been low at some point in the past); (iii) set 3, with codivergence set at -1 and all noncodivergence events at 0. Set 3 is a variation of set 1 with the same relative costs but where all costs are shifted to the left. This allows costs and codivergence events to be equated. Jane was run using the vertex-based cost mode, and the parameters of the genetic algorithm were kept at their default values (population size, 100; number of generations, 100). To determine the probability of observing the inferred costs by chance, costs were also calculated on a set of 500 samples for which tip mapping was randomized. Settings of the genetic algorithm were kept at default values.

Topology tests were performed to assess whether exceptions to a sce-

nario of perfect codivergence observed in the PyV phylogenetic tree were better supported by the data than a perfect codivergence model. This was done by using approximately unbiased (AU) tests, as implemented in CONSEL, version 0.1i (37).

Finally, divergence dates were also estimated. Topological congruence could emerge independently of codivergence, e.g., through preferential host switching (38). Observing synchronicity in timing of divergence events of hosts and their parasites reinforces the codivergence hypothesis. When viral lineage duplication occurs, synchronicity of parasite divergence events is also expected (provided the viral lineages maintain similar degrees of association to their host). Divergence date estimates were obtained using two methods: (i) as part of the aforementioned BMCMC analyses or (ii) by reestimating branch lengths of the ML tree under codon models using HyPhy, version 2.2.4 (39), and making the resulting tree ultrametric using a relaxed clock model implemented in r8s (40). The codon models used for this second set of analyses were a pure branch model derived of MG94, in which the ratio of nonsynonymous substitutions per nonsynonymous site to synonymous substitutions per synonymous site is estimated for each branch but assumed to be unchanged across sites (41), and an adaptive branch site random effects model (aBSREL), in which this ratio is estimated for each branch and allowed to vary across sites (42). We detected marked saturation at synonymous sites (data not shown); such strong saturation complicates analyses under both nucleotide and codon models. For both BMCMC and ML-based analyses, the relaxed clock was calibrated by setting a prior distribution (BEAST) or enforcing a fixed age (r8s) for the time to the most recent common ancestor (tMRCA) of lineage 1 using a published estimate of the split date of all hominine species (either 5.6 million years or a normal distribution with a mean of 5.6 million years and standard deviation of 0.25 million years [43]). Because we used the split date of all hominine species, estimates of times to the most recent common ancestors for viruses should be regarded as minimum bounds (viral coalescence times will necessarily predate the effective ancestral host population/species split). It should also be noted that divergence dates of the different hominine lineages are a point of active debate; this stems from both a scarce paleontological record and uncertainty in estimates of long-term mutation rates at genomic scales. For example, the estimate we opted for here (5.6 million years) is drawn from genomic analyses that proposed two estimates (5.6 or 11.2 million years), depending on priors on the substitution rates (1 \times 10⁻⁹ or 0.5 \times 10^{-9} mutations \cdot bp⁻¹ \cdot year⁻¹) (43). The focus of our synchronicity analyses was, however, on relative internode lengths, not absolute dates. Calendar years can thus be replaced with genetic distances and/or ratios of interest (see Table 6).

Accession numbers. Partial VP1 and whole-genome sequences, respectively, were deposited at the European Nucleotide Archive and GenBank under accession numbers LT158307 to LT158400 and KT184856 to KT184862.

Data availability. r8s and BEAUTi XML exemplary input files are available from the authors upon request.

RESULTS

Detection of short MCPyV-related sequences. Using a specific PCR system designed to amplify a ca. 700-bp fragment of the VP1 gene, we screened 386 fecal great ape and 197 human samples (Table 1). We detected MCPyV-related sequences in 50 great ape DNA extracts representing all hosts but Gorilla gorilla diehli, with fecal detection rates between 1.2% (Gorilla beringei beringei) and 53.8% (Pan paniscus). Nearly all sequences were found only at one site; a single sequence was detected in five and two Eastern chimpanzees (Pan troglodytes schweinfurthii) at two distinct sites in Uganda. For species/subspecies from which more than two sequences were obtained, considerable sequence divergence was observed; e.g., maximum observed distances were over 20%, possibly reflecting the circulation of viruses belonging to different lineages (discussed in more detail below). Minimum observed distances to publicly available sequences were often relatively high, i.e., between 5 and 17%. Finally, we also detected MCPyV sequences with >99% identity to published MCPyV sequences in 30 human DNA extracts (fecal detection rate, 15.2%). Most human DNA extracts were also screened with a PCR system intended to be lineage 1 specific (see below); all assays were negative.

Characterization of full genomes. We attempted to determine full-genome sequences from a selection of DNA extracts (n = 22). This was possible for samples from *P. paniscus* (n = 2), *P. troglo*dytes troglodytes (n = 3), P. troglodytes schweinfurthii (n = 1), and Gorilla beringei graueri (n = 1). Examination of putative open reading frames (ORFs) showed that all genomes displayed a typical PyV genome structure with an early region encoding regulatory proteins (small T and large T antigens) and a late region coding for structural proteins (VP1, VP2, and VP3) separated by a noncoding control region (NCCR). No open reading frame likely to encode a putative agnoprotein was identified. Overall, a ca. 80% sequence similarity to genomes of MCPyV and MCPyV-related nonhuman primate PyV was observed. Preliminary analyses revealed that the full genomes represented only a fraction of the overall genetic diversity detected in this study. To incorporate this broader diversity, we performed all of following phylogenetic analyses on an alignment of partial VP1 sequences (including sequences extracted from the novel full genomes).

Molecular phylogeny. We could not detect any signal indicative of recombination in the VP1 alignment (26). Phylogenetic analyses in both maximum likelihood (ML) (28) and Bayesian (31) frameworks supported the existence of a number of hostspecific clades (Fig. 1 and 2). All clades seemed to derive from three ancient lineages: one that comprised only MCPyV sequences and two that included only viral sequences detected in gorillas, bonobos, and chimpanzees. Branching order partially recapitulated host divergence events in the two great ape lineages (here, referred to as lineages 1 and 2) (Fig. 1 and 2). We identified four exceptions: (i) the polyphylies of PyV infecting Western chimpanzees in lineage 1 and Eastern chimpanzees in lineage 2, (ii) the interspersion of PyV infecting Eastern lowland and mountain gorillas in lineage 1, and (iii) the basal position of MCPyV.

Host specificity. We estimated the statistical support for host specificity using BaTS (Table 3). We found that viral sequences

found in a single host species were generally more likely to be closely related than expected by chance when both global and state-specific statistics were considered. The only exceptions corresponded to viral sequences identified in the sister subspecies *G. beringei beringei* and *G. beringei graueri*.

We also characterized the viral diversification process by running a species delineation analysis using general mixed Yule coalescent models (GMYC) (34, 35). The best GMYC model outperformed the null, full coalescent model (P = 0.0005) and identified 14 entities, among which 10 comprised several sequences. Nine multisequence entities comprised only sequences identified from a single host species/subspecies, indicating a close parallelism of PyV and host diversification processes (Fig. 1).

Codivergence. Taking the viral phylogeny presented in Fig. 1 as a given, we performed reconciliation analyses using Jane (Table 4). Under all tested cost sets, and whether the host species or subspecies phylogeny was considered, the number of codivergence events always exceeded the number of noncodivergence events. Randomization tests showed that, irrespective of the cost set, these results could not be explained by chance at the subspecies level. At the species level and using a P value threshold of 0.05, results obtained under two of the cost sets failed to reach statistical significance; it should, however, be noted that the species-level phylogeny comprises only five species, meaning that these tests had low power.

We also examined whether the viral topology presented in Fig. 1 was a better fit to our data than alternative topologies which enforced strict codivergence within lineages 1 and 2. The model forcing MCPyV to belong to lineage 1 was the only one that was rejected (AU test; P = 0.003). Monophyly of PyV infecting Western chimpanzees in lineage 1 and Eastern chimpanzees in lineage 2 as well as inclusion of MCPyV in lineage 2 could not be excluded (AU test; P = 0.52, 0.13, and 0.11, respectively). Given the very recent split of Eastern lowland and mountain gorillas (about 10,000 years ago [44]), the interspersion of PyV infecting these subspecies appeared biologically plausible, so we did not compare this scenario to a strict codivergence model.

Besides topological congruence, codivergence should result in synchronization of (i) viral and host divergence dates and (ii) viral divergence dates in the case of ancestral viral lineage duplication. We first estimated divergence dates using a relaxed clock model applied to nucleotide data in a Bayesian framework. For five of the six focal nodes of our analyses (nodes 1.2 to 4 and 2.1 to 3), these estimates were significantly older than host divergence events (Table 5). This pattern was compatible with the effects of the time dependency of molecular rates, i.e., the decay of molecular rates with increasing observation timescales, which can result in overestimating recent time to the most recent common ancestor (tMRCA) inferred from deep calibration points (19, 45-47). As this may arise through the effects of unaccounted-for purifying selection (among other possible mechanisms) (48, 49, 50), we reestimated all branch lengths using selection-aware models of codon evolution in an ML framework. A branch model of codon evolution resulted in divergence dates very close to those inferred by BMCMC analyses. Using an adaptive branch site random effect model of codon evolution, strong purifying selection was detected on a number of branches, including deep ones (data not shown). Most of the resulting increase in the overall tree length was supported by a single basal branch. This expansion prevented us from deriving any trustworthy tMRCA estimates.

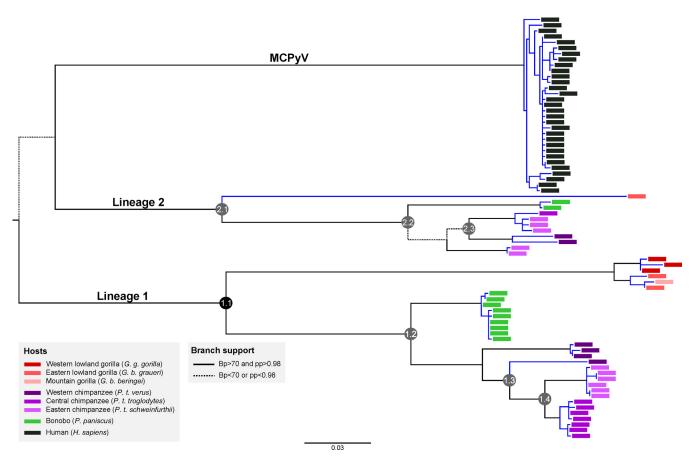


FIG 1 Maximum likelihood tree derived from an alignment of partial VP1 sequences. This tree was rooted at its center. The six gray circles stand for the main nodes whose date estimates are given in full in Tables 5 and 6; the black circle indicates the node that was used to calibrate the analyses. Note that these circles coincide with putative codivergence events. This tree was rooted using TempEst. Bp, bootstrap; pp, posterior probability. *G. g. gorilla*, *G. gorilla* gorilla; *G. b. graueri*, *G. beringei graueri*; *G. b. beringei*, *G. beringei*; *P. t. verus*, *P. troglodytes verus*; *P. t. troglodytes*, *P. troglodytes*; *P. t. schweinfurthii*, *P. troglodytes* schweinfurthii; *H. sapiens*, Homo sapiens.

Given the likely impact of strong purifying selection and our inability to properly account for it, we reexamined branch length/ internode ratios by rescaling the results in Table 5, using the tMRCA of a young node, node 1.4 (divergence of lineage 1 PyV infecting *P. troglodytes troglodytes* and *P. troglodytes schwein-furthii*), as a new arbitrary unit (Table 6). This resulted in a good agreement of host and virus relative divergence dates for most nodes (nodes 1.3 and 2.3 and nodes 1.2 and 2.2). The tMRCA of lineage 2 PyV infecting all great apes was a large underestimate of the divergence date of their hominine hosts, as expected under the hypothesis that deep branch lengths are severely underestimated.

DISCUSSION

The lack of any physical viral fossil record considerably complicates the task of understanding the long-term association of viruses with their hosts. However, using their present-day distribution, their nucleic acid sequences, and (more rarely) other biological traits, we can try to infer how long and how closely viruses have been associated with their hosts. The aim of this study was to determine whether codivergence, i.e., viral diversification driven by host diversification, is an important driver of PyV evolution.

Measurable host specificity is an absolute prerequisite for characterizing historical codivergence events. Host specificity has often been assumed for PyV, with only a few well-identified exceptions, e.g., budgerigar fledgling disease virus and simian virus 40 (SV40). Over the last decade, this assumption has been repeatedly supported by the implementation of generic PyV detection tools which have not revealed any multihost PyV (20, 51). Here, we used a PCR assay designed to specifically target a single PyV lineage to generate a large sample of sequences from closely related PyV infecting wild African great apes. Statistical tests strongly supported marked host specificity, which was still detectable at the host subspecies level. Viral diversification/speciation—as revealed by a GMYC model, i.e., according to the phylogenetic species concept—appeared strongly influenced by host diversification.

Host specificity and a coupling of viral diversification/speciation with host diversity could also arise over much shorter timescales than those implied by codivergence events. If codivergence is a dominant evolutionary process, a key expectation is that virus and host phylogenies should often be congruent. Phylogenetic analyses of great ape MCPyV-like sequences highlighted the existence of two viral lineages within which viral divergence events were mostly in line with hominine divergence events. Exceptions to the expectation of perfect codivergence within these lineages were not statistically supported. In addition, reconciliation analyses identified more codivergence events than noncodivergence

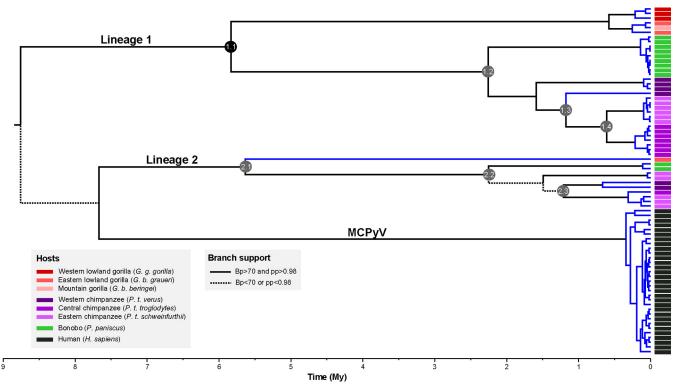


FIG 2 Chronogram derived from an alignment of partial VP1 sequences. This chronogram was obtained through BMCMC analyses run under a multispecies coalescent model (the clades corresponding to entities considered species are highlighted in blue). Other BMCMC analyses run under different tree priors and ML analyses gave similar results. The root of the tree was the most frequently observed in all posterior samples of trees (posterior probability of ca. 0.60) and was also retrieved by rooting the ML tree at its center. The six gray circles stand for the main nodes whose date estimates are given in full in Tables 5 and 6; the black circle indicates the node that was used to calibrate the analyses. Note that these circles coincide with putative codivergence events. Bp, bootstrap; pp, posterior probability.

Host species or subspecies	Mean association	Mean parsimony	Mean maximum exclusive single- state clade size ^b	Darahaa
(no. in group)	index	score	state clade size	P value
Species (5)	0.016	6		0
G. beringei			3	< 0.002
G. gorilla			3	< 0.002
H. sapiens			31	< 0.002
P. paniscus			9	< 0.002
P. troglodytes			17	< 0.002
Subspecies (8)	0.4	11		0
G. beringei beringei			1	1
G. beringei graueri			1	1
G. gorilla gorilla			3	< 0.002
H. sapiens			31	< 0.002
P. paniscus			9	< 0.002
P. troglodytes schweinfurthii			6	< 0.002
P. troglodytes troglodytes			7	< 0.002
P. troglodytes verus			3	< 0.002

TABLE 3 Results of BaTS tests for host specificity^a

^{*a*} The values reported are derived from analyses performed on posterior sets of trees generated under the 14-species coalescent model. Values were very similar when posterior samples of trees obtained under a constant population size coalescent model or a birth-death speciation model were analyzed.

^b Number of sequences.

events, irrespective of the host taxonomic level and cost set, e.g., 10 codivergence events versus 5 noncodivergence events considering host subspecies and all cost sets. Codivergence may therefore be the dominant process at play, accompanied by less frequent non-codivergence events, e.g., the viral lineage duplication event that gave rise to lineages 1 and 2.

On short timescales, host relatedness may influence viral transmission in such a way that topological congruence ensues in the absence of real codivergence, e.g., if host jumps are facilitated by host phylogenetic proximity (the preferential host switch hypothesis) (38, 52). A further step in validating codivergence events consists of showing that host and virus divergence events are synchronized. This requires branch lengths to be properly estimated

		No. of events ^a		
Host phylogeny	Cost set	Cospeciation	Not cospeciation	P value
Species level	1	5	2	0.056
	2	5	2	0.016
	3	5	2	0.066
Subspecies level	1	10	5	0
	2	10	5	0
	3	10	5	0

^a For the solution which was the most parsimonious in number of events.

TABLE 5 Absolute tin	TABLE 5 Absolute times to the most recent common ancestors	mon ancestor	s of PyV in lineages 1 and 2	1 and 2				
		Median tMF	RCA (10 ⁶ yr) or ML e	stimate (10 ⁶ yr; 95%]	Median t MRCA (10^6 yr) or ML estimate (10^6 yr; 95% HPD or Bp interval) b			
		Lineage 1				Lineage 2		
Statistical framework or reference	Diversification model or smoothing factor ^a	Node 1.1 (all) ^c	Node 1.2 (panine)	Node 1.3 (P. troglodytes)	Node 1.4 (P. troglodytes troglodytes + P. troglodytes schweinfurthii)	Node 2.1 (all)	Node 2.2 (panine)	Node 2.3 (P. troglodytes)
BMCMC	Coalescent, constant population size	5.62	2.15 (1.54–2.85)	1.09(0.74 - 1.48)	0.57 (0.36-0.85)	5.36 (3.71–7.31)	2.12 (1.47–2.92)	1.11 (0.75–1.55)
	Propromotion of the second second Speciation, birth-death	5.62 5.62	2.25 (1.57–3.11) 2.06 (1.46–2.72)	1.18 (0.79–1.67) 1.05 (0.71–1.44)	$0.61 (0.37-0.91) \\ 0.54 (0.34-0.79)$	5.63 (3.84–7.95) 5.27 (3.58–7.20)	2.25 (1.52–3.18) 2.04 (1.38–2.80)	1.21 (0.79–1.73) 1.07 (0.69–1.48)
ML	1 100	5.62 5.62	2.29 (1.61–4.56) 2.24 (1.61–4.19)	$1.06 (0.76 - 2.22) \\ 1.04 (0.70 - 2.27)$	$0.54 (0.23-0.84) \\ 0.53 (0.18-1.93)$	5.51 (3.97–24.92) 5.54 (4.02–9.84)	2.27 (1.77–25.37) 2.26 (1.79–4.98)	1.21 (0.75–13.33) 1.21 (0.76–2.24)
Prado-Martinez et al. (43) ^d		5.62	0.87	0.42	0.17	5.6	0.87	0.42
^{<i>a</i>} Diversification models we ^{<i>b</i>} Estimates that are incomp Bootstrap intervals were de	re used for Bayesian Markov cha atible with those determined in termined using 100 bootstrap ps	ain Monte Carlo Prado-Martinez seudoreplicates o	(BMCMC) analyses; sm et al. (43) appear in bol of the codon data set fror	toothing factors were use d. The 95% highest poste n which branch lengths v	^a Diversification models were used for Bayesian Markov chain Monte Carlo (BMCMC) analyses; smoothing factors were used for maximum likelihood (ML) analyses (under the MG94-like model of codon evolution). ^b Estimates that are incompatible with those determined in Prado-Martinez et al. (43) appear in bold. The 95% highest posterior density (HPD) was used for BMCMC analyses; bootstrap (Bp) intervals were used for ML analyses. Bootstrap intervals were determined using 100 bootstrap pseudoreplicates of the codon data set from which branch lengths were restimated on the ML topology; all trees were rooted using TempEst. tMRCA, time to the most recent	alyses (under the MG94-l ACMC analyses; bootstrap y; all trees were rooted usi	ike model of codon evolui (Bp) intervals were used ing TempEst. tMRCA, tim	ion). for ML analyses. e to the most recent

throughout the phylogeny. Here, we speculate that the well-documented time dependency of molecular rates, which posits an apparent decay of molecular rates with increasing measurement timescales (19, 45–47), may have resulted in overestimating recent divergence dates derived from our initial molecular clock analyses which were calibrated with an ancient divergence event. In line with this hypothesis, we found that the relative timescales of host and virus divergence events were in good agreement when these estimates were rescaled using an arbitrary unit set to a recent divergence event, i.e., a procedure similar to calibrating the molecular clock with this recent divergence event. In addition, codivergence events were also synchronous in the viral lineages 1 and 2.

Overall, we observed (i) marked host specificity, (ii) frequent codivergence events, and (iii) the synchronicity of a number of codivergence events. The evolution of MCPyV-related viruses with their hominine hosts therefore appears to have been mostly driven by host-PyV codivergence. A number of other human PyV have been shown to be closely related to great ape PyV (22, 53–56). The respective lineages may represent promising opportunities to test whether the dominance of codivergence events can be generalized throughout the PyV family tree. Regardless, the findings reported here lend support to the hypothesis of an ancient association of PvV and their animal hosts, which the well-known separation of mammal and bird PyV and the recent discovery of the first fish and arthropod PyV already pinpointed (6, 8, 57). In a recently published phylogeny based on the large T antigen, the root age of the family tree was more than 11 times the age of the MRCA of MCPyV-related viruses (20). Assuming this MRCA dates back to about 6 million years ago, the family root would be more than 60 million years old. Assuming that the PyV family tree is affected by the phenomenon of time dependency of molecular rates, the root age of the family may be even more ancient, as recently suggested by C. B. Buck et al. (8).

Although a robust signal for codivergence exists, we did not observe strict codivergence of MCPyV-related viruses and their hominine hosts. For example, in our phylogenetic analyses, the placement of the MCPyV lineage is ambiguous, and the most ancient divergence event of polyomaviruses apparently postdates the respective divergence event of their hominine hosts. Although these observations may be explained by limitations of the models of sequence evolution we used, we cannot exclude the hypothesis that they reflect biological reality. Since hominine species are recently diverged, the combination of ancestral viral diversity and incomplete lineage sorting may suffice to explain apparent deviations from strict codivergence; i.e., perfect patterns of codivergence are not necessarily expected, even where no other processes have been at play (19). However, a notion emerging in the literature is that a mixture of processes, including, but not restricted to, measurable codivergence with their hosts, will generally provide a better explanation for ds-DNA virus evolution in the long run than strict codivergence. For example, it was proposed that herpes simplex virus 2 (HSV-2) arose as a consequence of the transmission of a chimpanzee simplex virus to the human lineage (50). Similarly, host switches as well as lineage duplications have been documented in papillomaviruses (2, 3). It seems clear that processes other than codivergence were also at play during PyV evolution, as notably illustrated by the 13 human PyV identified thus far and

le-9 mutations/base pair/year.

Assuming a mutation rate of

The respective node

common ancestor

was used to calibrate the trees.

		Median tM	RCA or ML estimate	Median tMRCA or ML estimate (95% HPD or Bp interval) b	erval) ^b			
		Lineage 1				Lineage 2		
Statistical framework or reference	Diversification model or smoothing factor ^a	Node 1.1 (all) ^c	Node 1.2 (panine)	Node 1.3 (P. troglodytes)	Node 1.4 (P. troglodytes troglodytes + P. troglodytes schweinfurthii)	Node 2.1 (all)	Node 2.2 (panine)	Node 2.3 (P. troglodytes)
BMCMC	Coalescent, constant	9.82	3.77 (2.70-5.00)	1.91 (1.30-2.60)	1 (0.63–1.49)	9.40 (6.51–12.82)	3.72 (2.58-5.12)	1.95 (1.31-2.72)
	population size Multispecies coalescent	9.18	3.69(2.57 - 5.10)	1.93 (1.29–2.74)	1 (0.61 - 1.49)	9.23 (6.29–13.03)	3.69 (2.49–5.21)	1.98 (1.29–2.84)
	Speciation, birth-death	10.37	3.81 (2.70-5.04)	1.94 (1.31-2.67)	1 (0.63–1.46)	9.76 (6.63–13.33)	3.78 (2.56-5.18)	1.98 (1.28–2.74)
ML	1	10.37	4.24(2.98 - 8.44)	1.96(1.41 - 4.11)	1 (0.43–1.58)	10.20 (7.35-46.15)	4.20 (3.28-46.98)	2.24 (1.39-24.68)
	100	10.60	4.22 (3.04-7.90)	1.96 (1.32–4.28)	1 (0.34 - 3.64)	10.45 (7.58–18.57)	4.26 (3.38–9.40)	2.28 (1.43-4.22)
Prado-Martinez et al. (43)		32.11	4.98	2.39	1	32.11	4.98	2.39
" Diversification models we ^b Estimates that are incomp The 95% highest posterior branch lengths were reestin	rre used for Bayesian Markov cl patible with those determined in density (HPD) was used for BM nated on the ML topology: all t	hain Monte Carle n Prado-Martine: 1/CMC analyses; 1/CMC analyses	o (BMCMC) analyses; sr z et al. (43) appear in bo bootstrap (Bp) intervals using TempEst	noothing factors were us ld. For time to the most were used for ML analys	^{<i>a</i>} Diversification models were used for Bayesian Markov chain Monte Carlo (BMCMC) analyses; smoothing factors were used for maximum likelihood (ML) analyses (under the MG94-like model of codon evolution). ^{<i>b</i>} Estimates that are incompatible with those determined in Prado-Martinez et al. (43) appear in bold. For time to the most recent common ancestor (tMRCA), 1 unit = tMRCA (<i>P. troglodytes troglodytes + P. troglodytes schweinfurthii</i>). ^{<i>b</i>} The 95% highest posterior density (HPD) was used for BMCMC analyses; bootstrap (Bp) intervals were used for ML analyses. Bootstrap intervals were determined using 100 bootstrap pseudoreplicates of the codon data set from which branch lengths were restimated on the ML topology: all trees were noted using TempEst.	nalyses (under the MG94-li 1 unit = tMRCA (<i>P. troglo</i> ined using 100 bootstrap ps	ke model of codon evolut <i>dytes troglodytes</i> + <i>P. trog</i> eudoreplicates of the codo	ion). <i>lodytes schweinfurthi</i> i). 2n data set from which
branch lengths were reestir ^c No values for highest post	^c No values for highest posterior density or bootstrap interval because this node was used to calibrate the trees	rees were rooted val because this 1	using TempEst. node was used to calibra	te the trees.				

the two great ape lineages documented in this study. Further biological characterization of representatives of these lineages may reveal whether these noncodivergence events were driven by adaptive change, e.g., tissue tropism, or stochastic, e.g., demographic, processes (58).

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